

1 **Dendroecology in common gardens: Population differentiation and plasticity in**
2 **resistance, recovery and resilience to extreme drought events in *Pinus pinaster***

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4 *Running title: Maritime pine resilience to extreme droughts*

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29 **Abstract**

30 Quantifying intraspecific genetic variation and phenotypic plasticity of traits involved in
31 drought tolerance is essential to forecast forest tree vulnerability to climate change. Tree
32 ring analysis was applied to retrospectively assess the resistance, recovery and resilience
33 of a Mediterranean pine in the face of extreme climate episodes. We combined a
34 dendrochronological approach with the analysis of common-garden tests to disentangle
35 genetic, environment and genetic-by-environment effects in growth responses of 10 *P.*
36 *pinaster* populations to two extreme climatic events. Trees were 43 years old and had
37 experienced two intense droughts, in 1995 and 2005. Drought events caused drastic
38 reductions in secondary growth, but trees showed high capacity to recover pre-drought
39 growth rates. The differences in the characteristics of the two drought events and the
40 environmental distance between sites strongly modulated maritime pine responses to
41 extreme droughts. However, a common among-population signal across sites and events
42 was detected in the drought response strategy. Among-population variation in response
43 to extreme droughts was evident for the resistance and recovery components, two
44 strategies that appeared to trade-off between each other. Populations from Atlantic
45 climates showed higher resistance but lower recovery capacity, whereas Mediterranean
46 origins prioritised recovery over resistance. Mediterranean populations showed a more
47 conservative strategy that indicated an adaptive advantage under water stress, reflected
48 in greater long-term survival. The abovementioned relationships were clear in the site
49 where the impact of the drought events was strongest, but not in the more favourable
50 site. Differences in relationships between sites reflect that strategies of populations to
51 cope with drought are strongly context dependent. Based on these results, we infer that
52 future extreme droughts will differentially affect *P. pinaster* populations across the
53 natural range of the species. Immediate effects will be more evident in Mediterranean
54 areas but, in the long term, population persistence in the face of climate change will be
55 more compromised for Atlantic origins. Because local environmental conditions can
56 considerably modulate responses to extreme events, special attention is required to
57 define appropriate management practices to mitigate the impact of future droughts.

58

59 **Key words:** Adaptive variation, Common garden, Climate change, Dendroecology,
60 Drought tolerance, Extreme drought events, Phenotypic plasticity, Recovery,
61 Resilience, Resistance

62

63 **Introduction**

64 Ongoing climate change is imposing serious constraints on forest productivity and
65 persistence worldwide (Parmesan, 2006). The increasing frequency and intensity of
66 extreme climate events such as droughts, heat waves and storms drastically impact tree
67 growth and trigger forest dieback (Allen et al., 2010; Neumann et al., 2017). Extreme
68 droughts are notably relevant in Mediterranean climate areas, where evidence is
69 accumulating of forest decline due to drought (Camarero et al., 2015; Martínez-Vilalta
70 and Piñol, 2002; Sánchez-Salguero et al., 2010).

71 Forest persistence in the face of climate disturbance may depend on the ability of
72 trees to tolerate and recover from disturbance and maintain their function during and
73 after the event (Ingrisch and Bahn, 2018). Drought tolerance relies on a wide range of
74 physiological plastic responses to water deficit (McDowell et al., 2008; Ryan, 2011). In
75 conifers, these physiological responses include mechanisms intended to (i) avoid
76 hydraulic failure by stomata closure regulation, (ii) prevent cell damage under low
77 water potential by osmotic adjustments, (iii) prevent carbon starvation by reallocating
78 carbon reserves, and (iv) improve water uptake and use efficiency by changes in growth
79 allocation priorities (Moran et al., 2017). Individual plastic responses may, however, be
80 insufficient to cope with severe, prolonged or concatenated droughts, resulting in
81 increased vulnerability to drought and high mortality rates (Navarro-Cerrillo et al.,
82 2018). In these cases, tree populations can persist only by adaptive genetic changes.
83 Selection pressure on adaptive traits related to drought tolerance may result in long-term
84 changes in the mean and variance of population tolerance to drought. Gene flow from
85 drought-adapted genetic pools may also help to enhance the drought tolerance of a
86 population. However, migration and natural selection operate more slowly than the

87 current pace of climate change, compromising forest persistence in the face of climate
88 disturbances (Jezkova and Wiens, 2016).

89 Considerable effort has been made in recent decades to determine the
90 quantitative genetic basis of drought adaptation. Most studies have focused on
91 manipulative experiments on young trees (Moran et al., 2017), reporting large among-
92 and within-population variation in drought tolerance (e.g. de la Mata et al., 2014;
93 Gaspar et al., 2013; Lamy et al., 2014). Particularly, populations of maritime pine
94 (*Pinus pinaster* Ait.) from dry climates were reported to be adapted to drought through
95 larger biomass allocation to roots (Corcuera et al., 2012), higher water use efficiency
96 with lower dependence on stomata closure regulation (Correia et al., 2008), higher
97 resistance to cavitation (Corcuera et al., 2011; but see Lamy et al., 2014), and higher
98 osmotic adjustment (Nguyen-Queyrens and Bouchet-Lannat, 2003) than populations
99 from mild climates. Conversely, populations from mild climates were reported to grow
100 faster when conditions were favourable and exhibited a less conservative strategy
101 against abiotic stress (Corcuera et al., 2012; de la Mata et al., 2014). Tolerance at young
102 ages is, however, not necessarily transferable to older ages (Cavender-Bares and
103 Bazzaz, 2000; He et al., 2005). Because drought is also an important selective agent in
104 mature stages, either directly, by causing hydraulic failure (Martínez-Vilalta and Piñol,
105 2002), or indirectly, by contributing to interactions with biotic stress (Netherer et al.,
106 2015), quantitative analysis of drought tolerance genetics in mature trees is required.

107 Dendrochronological approaches have recently been applied to examine climate
108 sensitivity and responses to extreme drought events in mature trees (Gazol et al., 2017;
109 Gazol et al., 2018; Rubio-Cuadrado et al., 2018; Sánchez-Salguero et al., 2018; Serra-
110 Maluquer et al., 2018). Dendrochronology provides an annually-resolved record of tree
111 responses to climate variation across the tree lifespan. This technique can be applied to

112 assess the impact on trees and their recovery capacity in the face of extreme climate
113 events by analysing annual radial growth before, during and after a single extreme
114 climate event (Lloret et al., 2011). Recent studies have shown that the ability of trees to
115 cope with extreme climate events varies greatly among species (e.g. Gazol et al., 2018)
116 and across the natural range of a given species (e.g. Sánchez-Salguero et al., 2018).
117 Specifically, maritime pine populations growing in dry sites were less resistant to
118 drought but recovered faster than trees growing in mesic sites (Sánchez-Salguero et al.,
119 2018). However, these studies were not able to properly disentangle environmental from
120 genetic effects as they were based on surveys of natural populations. Disentangling
121 plastic from genetic variation is mandatory to properly assess the impact of climate
122 change across the distribution range of a species. This can be achieved only by
123 quantitative genetic approaches in which different populations are grown together in a
124 common environment (Montwe et al., 2016).

125 Several studies have analysed intraspecific variation in resilience to extreme
126 drought events in common garden tests using dendrochronological approaches (George
127 et al., 2017; George et al., 2015; Heer et al., 2018; Housset et al., 2018; Montwe et al.,
128 2016; Montwé et al., 2015; Taeger et al., 2013; Trujillo-Moya et al., 2018). These
129 studies suggested that patterns of within-species variation in strategies to cope with
130 extreme droughts differ greatly depending on the species considered, as has been
131 previously observed in across-species surveys in natural systems (Gazol et al., 2018).
132 Most of the previous studies focused on a single extreme event (Montwe et al., 2016) or
133 a single test site (Housset et al., 2018), but responses to extreme events are dependent
134 on the event itself (Gazol et al., 2018; Serra-Maluquer et al., 2018) and the
135 environmental conditions of the test site (Savva et al., 2007). Moreover, the among-
136 population patterns of variation in the response to extreme events can be affected by

137 variable plastic responses to the environment of the test site (Savva et al., 2007). To
138 assess the potential of local adaptation of tree species or populations to climate change,
139 it is essential to understand the stability of the strategies and components of genetic
140 variation in drought tolerance traits across a range of extreme climate events and across
141 sites.

142 We analysed the influence of genetic (G, i.e. population differentiation),
143 environment (E, site-to-site plasticity) and G×E interaction (i.e., among-population
144 variation in phenotypic plasticity) in the growth responses of maritime pine (*Pinus*
145 *pinaster* Ait.) populations to two episodes of extreme drought using a
146 dendrochronological approach. Annual increment in basal area in the last 30 years was
147 quantified in adult pine trees from 10 populations covering most of the distribution
148 range of the species, planted in two replicated common garden tests in central Spain.
149 Resistance (i.e. the impact of an extreme event on current growth relative to previous
150 growth), recovery (i.e. the relative increase of growth rates after the event) and
151 resilience (i.e. the capacity to reach pre-episode growth levels) in the face of the two
152 extreme drought events identified were estimated following Lloret et al. (2011).
153 Quantitative genetic analysis allowed us to estimate population differentiation across
154 sites and events in these resilience components and identify signs of adaptation
155 associated with the environmental conditions at the origin of each population.

156

157 **Material and Methods**

158 *Study system*

159 Maritime pine (*Pinus pinaster* Ait.) is an important Mediterranean tree that covers more
160 than 4 Mha in Southwest Europe and North Africa. Across its natural range, this species
161 grows in variable environmental conditions, from mild environments in the Atlantic to

162 harsh, dry Mediterranean climates, and from sea level to more than 2000 m in the
163 southern regions. Its natural distribution is highly fragmented and, either because of
164 neutral demographic processes or local adaptation, populations are strongly
165 differentiated in many life history traits, including primary (height) and secondary
166 (radial) growth (Benito-Garzon et al., 2013; Di Matteo and Voltas, 2016) and tolerance
167 to abiotic stress (Chambel et al., 2007; de la Mata et al., 2014; but see Lamy et al.,
168 2014). Radial growth of mature maritime pine trees was shown to be sensitive to water
169 availability (Arzac et al., 2018; Bogino and Bravo, 2008), and this effect varied among
170 populations depending on their local environmental conditions (Caminero et al., 2018;
171 Rozas et al., 2011).

172

173 *Plant material and common garden experiments*

174 The present study is based on data collected from two provenance trials established in
175 1967 as part of a larger series of test sites where 52 maritime pine populations from the
176 complete natural range of the species were planted in five different sites in central Spain
177 (Alia et al., 1997; Benito-Garzon et al., 2013). Each trial shared the same populations,
178 and each population was represented by seed lots from 30 unrelated mother trees. The
179 experimental design in each test site comprised a complete block design with four
180 blocks of experimental units of 16 trees per population, planted at 2.5×2.5 m spacing.

181 We focused on the Cabañeros and Riofrío sites, in Ciudad Real province (Fig.
182 1a) and on 10 contrasting populations belonging to seven of the eight genetic pools
183 identified by neutral markers for this species (Bucci et al., 2007) (Table ST1). The
184 populations were selected to cover as much of the environmental variation across the
185 natural distribution range of maritime pine as possible (Fig. 1b, Table ST1). The
186 conditions at origin for these populations were highly variable, ranging from 28 to 1600

187 m in elevation, 336 to 1555 mm in total annual precipitation, and 9.7 to 15.6 °C in mean
188 annual temperature (Table ST1).

189 The two test sites are located in central Spain under Mediterranean continental
190 climate, although Cabañeros is slightly wetter and colder than Riofrío (Table ST1). The
191 two sites have acidic soils (pH ~ 5) derived from quartzite bedrock, but soils in Riofrío
192 are stonier and shallower than in Cabañeros. Additionally, Cabañeros is on a
193 homogeneous hilltop plain while Riofrío is on a mid-hill with irregular topography.
194 Altogether, although the environmental differences between the two sites are not that
195 big (Fig. 1b), they were enough to imprint notable differences in pine performance, with
196 higher growth and survival rates of pines in Cabañeros than in Riofrío (Table ST1).

197 For both sites, monthly weather records from tree establishment to sampling was
198 estimated using data from nearby meteorological stations of the National Agency of
199 Meteorology (AEMET, Spanish Government; see details in Supplementary Information
200 Methods SM1). Standardised precipitation-evapotranspiration index (SPEI) was
201 calculated from the monthly temperature and precipitation series using the *SPEI*
202 package of the software R 3.5.2 (Vicente-Serrano et al., 2010). Extreme drought events
203 in the last three decades at both test sites were identified as minimum values of mean
204 SPEI from October to May, the rainy period in the central Iberian Peninsula. Two
205 extreme drought events were identified in 1995 and 2005 at both sites (Fig. 2a). These
206 events were identified and used in previous studies that analysed drought responses in
207 central and southern Spain (Rubio-Cuadrado et al., 2018; Sánchez-Salguero et al.,
208 2010).

209

210

211

212 *Climate data*

213 Monthly mean climate data at each population origin were obtained for the period 1950-
214 2000 from a regional climate model (Gonzalo-Jimenez, 2008) in the case of the seven
215 Spanish populations, and from the Climate Explorer of the Royal Netherlands
216 Meteorological Institute (CRU TS 4.01 data set; <http://climexp.knmi.nl>; Mitchell and
217 Jones (2005)) for the French, Italian and Moroccan populations. The model of Gonzalo-
218 Jimenez (2008) is known to be more accurate in the Iberian Peninsula than world-level
219 models because it takes into consideration a dense network of meteorological stations
220 (Jaramillo-Correa et al., 2015). Both models have a resolution of 1 x 1 km; for each
221 population, data from the corresponding pixel was retrieved. The climate variables were
222 annual mean temperature, maximum temperature in the warmest month, minimum
223 temperature in the coldest month, temperature seasonality, annual and summer
224 precipitations, and average number of frost days per year. To minimise Type I error in
225 the correlation analyses between population performance and climate (see below), the
226 information from these seven climate variables was summarised into two main
227 components through a PCA using PROC PRINCOMP in SAS. The two first PCs
228 (hereafter 'climate indices') explained ~84% of the overall variation. Climate index 1
229 explained 46.5% of the variance and was positively related to high minimum
230 temperatures and precipitation and inversely related to temperature seasonality,
231 suggesting a proxy of Atlantic climate for higher values of the index and of Continental
232 and Mediterranean climates for lower values of the index (Fig. 1b). Climate Index 2
233 explained 37.2% of the variance and was positively related to mean temperatures and
234 negatively related to frost frequency, suggesting a thermal gradient from cold to warm
235 conditions (Fig. 1b).

236

237

238 *Sampling and tree-ring data processing*

239 In summer 2011, when trees were 43 years old, tree survival of the populations was
240 assessed and wood samples were taken from 12-18 trees per population (3-5 trees per
241 block) at each test site. Trees were cored at 1.3 m above ground using 5-mm Pressler
242 increment borers. Two to four wood cores were sampled from each tree. Wood cores
243 were sanded until tree rings were clearly visible and tree-ring series were visually cross-
244 dated by assigning calendar years to the rings through identification of characteristic
245 wide and narrow ring sequences (Speer, 2010). Tree-ring widths were measured to the
246 nearest 0.001 mm on 2–4 radii per tree using a binocular microscope and a linear stage
247 (Velmex Inc., Bloomfield NY, USA) interfaced with a computer. The accuracy of cross
248 dating and ring width measurement was checked using the COFECHA programme
249 (Holmes, 1983).

250 After averaging all ring width series from each tree on an annual basis, individual
251 basal area increment (BAI, $\text{cm}^2 \text{yr}^{-1}$) series were calculated from ring widths, assuming
252 circularity of the rings section and considering all available rings in each core (Biondi
253 and Qeadan, 2008). BAI chronologies for each population and test site were calculated
254 as a year-by-year arithmetic mean of individual BAI series. To avoid missing segments
255 in the BAI series at young ages, the 31-year common period to all series from 1980 to
256 2010 was considered for the analyses.

257

258 *Pointer years*

259 To assess the impact of drought events on growth stability, we calculated the negative
260 pointer years, i.e., years with extremely low annual growth in relation to previous years
261 in a high proportion of trees (Schweingruber et al., 1990), using the *dplR* package

262 (Bunn, 2010). Negative pointer years were calculated considering BAI data for each
263 individual tree at each test site. When more than 50% of trees at a site exceeded 50%
264 BAI reduction in comparison to mean BAI for the previous three years, the year was
265 considered a negative pointer year.

266

267 *Drought resilience components*

268 To quantify tree growth response to extremely dry conditions, we calculated the four
269 resilience components proposed by Lloret et al. (2011): resistance (Rt), recovery (Rv),
270 resilience (Rs) and relative resilience (rRs), calculated from BAI data as:

$$271 \quad R_t = \frac{BAI_{DY}}{BAI_{PreDY}}$$

$$272 \quad R_v = \frac{BAI_{PostDY}}{BAI_{DY}}$$

$$273 \quad R_s = \frac{BAI_{PostDY}}{BAI_{PreDY}}$$

$$274 \quad rRs = \frac{BAI_{PostDY} - BAI_{DY}}{BAI_{PreDY}}$$

275 where BAI_{DY} is the BAI during the drought year, BAI_{PreDY} is the mean BAI during the
276 three pre-drought years, and BAI_{PostDY} is the mean BAI during the three post-drought
277 years. R_t quantifies the capacity of trees to buffer drought stress and continue growing
278 during drought. R_v quantifies the growth reaction during the three years following the
279 drought year. As recommended, R_v was not calculated when BAI_{DY} was zero (Lloret et
280 al., 2011). R_s quantifies the capacity of trees to recover pre-drought growth rates, and
281 rRs quantifies the net balance between growth during extreme drought and the three
282 following years. As the different indices share some elements in their equations, it
283 should be noted that they are not completely independent (e.g. resistance and recovery
284 could tend to be negatively related).

285

286 *Inter-tree competition*

287 As response to drought events may be contingent on the amount of resources each tree
288 has available (Martinez-Vilalta et al., 2012), we accounted for competition interference
289 in our statistical models. For that purpose, we assessed the surrounding basal area of
290 each tree before each climatic event as a proxy of neighbouring competition (Serra-
291 Maluquer et al., 2018). To estimate this competitive index, we took advantage of tree
292 survival and dendrometric assessments done in 1985 (Alia et al., 1995) and 1998 (Alía
293 et al., 2001). To homogenize the interval between the previous assessments and the two
294 drought climatic events (occurring in 1995 and 2005), tree diameter in 1995 was
295 interpolated from the assessments done in 1985 and 1998, assuming a constant growth
296 rate during this period. Basal area surrounding each core-sampled tree was then
297 estimated by adding the basal area at breast height of the immediately surrounding trees
298 that were alive (8 neighbours at most) ten years before each climatic event.

299

300 *Statistical analyses*

301 A repeated mixed model analysis applied to tree-level BAI data was used to compare
302 mean BAI before (three years before), during (event year) and after (three years after)
303 the drought event and to explore whether differences among these three periods
304 depended on the test site, the pine population and their interaction (see details in
305 Methods SM2). In these models, the drought factor (three levels: before, during, after)
306 accounts for the plastic response in BAI to the drought event (i.e. temporal plasticity
307 within each subject) while the interactions between this factor and the sites and the
308 populations represent spatial plasticity to environmental variation between sites and
309 among-population variation in this spatial plasticity, respectively.

310 Resilience components were analysed by repeated measures mixed models but,
311 in this case, the models were fitted for the two drought events (1995 and 2005) together.
312 Individual trees were the subjects for which two repeated measures (one for each event)
313 were obtained. The events and their interaction with sites and populations were thus
314 considered within-subject repeated measures. The mixed model used was (random
315 effects in italics):

316

$$317 \text{ RC} = \mu + \text{S} + \text{POP} + \text{POP} \times \text{S} + \text{Ev} + \text{Ev} \times \text{S} + \text{Ev} \times \text{POP} + \text{Ev} \times \text{POP} \times \text{S} + \text{B}(\text{S}) +$$
$$318 \text{B} \times \text{POP}(\text{S}) + \text{Tree}(\text{B} \times \text{POP} \times \text{S}) + \text{NBA} + \varepsilon$$

319

320 where RC is the resilience component of each tree and event, μ is the overall mean; S,
321 POP, and POP \times S are across-subject fixed effects that account for global common
322 differences in the resilience component between sites, populations and their interaction;
323 Ev is the within-subject fixed effect (repeated measures on the same trees) of the
324 extreme event (1995, 2005); B(S) and B \times POP(S) represent the random variation
325 between the blocks within sites and between the whole plots of the multi-tree block
326 design, respectively; *Tree*(B \times POP \times S) is a random effect identifying each tree that
327 account for the autocorrelation among the repeated measures within trees; NBA is a
328 continuous fixed covariate that accounts for neighbouring basal area, and ε is the
329 random residual. The event factor indicates differences in the resilience component
330 between the two events, while the Ev \times S, Ev \times POP and Ev \times POP \times S interactions reveal
331 whether variation in the resilience component between sites and populations is
332 contingent on each particular event. Covariation with neighbouring basal area accounts
333 for the modulation of the resilience components by competitive effects. Mixed models
334 were fitted with the MIXED procedure of the SAS System (Littell et al., 2006).

335 To determine whether variation among populations in response to drought
336 events was related to climate conditions at the origin of the populations, across-
337 population Pearson correlations (CORR procedure in SAS) were carried out between
338 the resilience components and the two climate indices obtained from the principal
339 component analysis described previously. Relationships between the different resilience
340 components at the population level were also tested with Pearson correlation analysis.
341 Correlation analyses were conducted using the across-extreme-events least square
342 means of the resilience components for each population at each test site as derived from
343 the previous mixed models.

344

345 **Results**

346 Standardised precipitation-evapotranspiration indices from previous October to current
347 May varied greatly from year to year across the study period (1980-2010), and two
348 extreme drought episodes in 1995 and 2005 at both the Cabañeros and Riofrío test sites
349 were identified (Fig. 2a). In Cabañeros, considerable reduction in radial growth rates of
350 more than 50% of sampled trees (i.e., negative pointer year) was observed only in 2005
351 (Fig. 2b). However, in Riofrío, both dry years coincided with negative pointer years
352 (Fig. 2c).

353

354 *Growth before, during and after the drought events*

355 Basal area increment (BAI) drastically declined during the extreme drought events but
356 recovered afterwards (Fig. 3). Growth responses to the drought events were, however,
357 contingent on the site (Fig. 3) and the population (Fig. SF1). Reductions of BAI during
358 the two extreme drought events were more pronounced in Riofrío than in Cabañeros,
359 but growth recovery after the events was stronger in Riofrío than in Cabañeros (Fig. 3).

360 For both extreme events, all populations showed a drastic decrease in radial growth
361 during the drought year but later recovered normal pre-drought growth rates (Fig. SF1).
362 However, growth rates of the PRAV, COCA and TAMJ populations were significantly
363 higher after the 2005 drought event than before (Fig. SF1).

364

365

366 *Variation in resilience components*

367 Resistance, recovery and relative resilience of growth to the drought events were
368 significantly influenced by the sum of basal areas of neighbouring trees, a proxy of
369 inter-tree competition (Table 1). Trees subjected to higher inter-tree competition
370 showed reduced resistance to the extreme events, but higher recovery afterwards as well
371 as higher relative resilience. No effect of competition on drought resilience was
372 observed (Table 1).

373 Large differences were observed in the resilience components between the two
374 extreme events (event factor in Table 1). Significant variation between sites and among
375 populations was also observed across events for most of these components (site and
376 population factors in Table 1). Site differences were particularly high for resistance,
377 recovery and relative resilience (Table 1), with trees in Cabañeros consistently showing
378 higher resistance but lower recovery and relative resilience than trees in Riofrío for the
379 two events (Fig. 4). Differences between sites in resistance and relative resilience were,
380 however, greater for the 1995 event than for the 2005 event (Fig. 4). Population
381 resilience did not differ between sites for either event (Table 1, Fig. 4).

382 After accounting for the potential bias generated by variable competitive effects,
383 resistance to and recovery from drought events also varied significantly among pine
384 populations, although this variation differed between events or sites (see interaction

385 terms with the population factor in Table 1). Despite these interactions, both population
386 resistance and population recovery were positively correlated between events in Riofrío
387 ($r = 0.89$, $N = 10$, $p < 0.001$ for resistance, and $r = 0.89$, $N = 10$, $p < 0.001$ for recovery)
388 but not in Cabañeros (Fig 4a,b). In general, populations from the Atlantic climate (e.g.,
389 LEIR, PRAV, CAMB) showed higher resistance but lower recovery than populations
390 from drier conditions (e.g., CARA, TAMJ) (Fig. 4). Population resistance across the
391 two events was negatively related to population recovery, although the relationship was
392 significant in Riofrío ($r = -0.73$, $N = 10$, $p = 0.018$) and not in Cabañeros ($r = -0.37$, N
393 $= 10$, $p = 0.286$), where the recovery values and their among-population variation were
394 very low (Fig. SF2). It should be noted that the CARA population, from southeast
395 Spain, showed outstanding capacity for growth recovery after drought events, and this
396 outstanding recovery may be forcing the previously reported relationship. However, the
397 relationship between resistance and recovery in Riofrío remained significant after
398 removing this population from the analysis ($r = 0.76$, $N = 9$, $p = 0.017$).

399 Among-population variation in resilience and relative resilience was site- and/or
400 event-dependent (population \times site and population \times event interactions, respectively, in
401 Table 1), resulting in no overall differences among populations in these parameters.

402

403 *Dependence of drought resilience components on climate conditions at origin*

404 Population resistance to the extreme drought events in Riofrío was significantly and
405 positively related to climate index 1 (Fig. 5a), a proxy of Atlantic conditions at positive
406 values of the index and of continental Mediterranean conditions at negative values.
407 Growth resilience in Cabañeros was also positively related to climate index 1 (Fig. 5c),
408 although this relation should be interpreted with caution because no significant overall
409 variation among populations or population \times site interactions were detected for this trait

410 (Table 1). No other significant relations were detected between resilience components
411 and climate conditions at population origin (Fig. 5) except for a negative correlation ($r =$
412 -0.74 , $N = 10$, $p = 0.013$) between growth recovery in Cabañeros and climate index 2, a
413 proxy of thermal conditions (Fig. SF3).

414 Interestingly, among-population variation in growth resistance to drought events
415 in Riofrío was negatively correlated with accumulated survival at age 43 (Fig. 6a).
416 Population survival was also positively related to recovery from drought events in
417 Riofrío, although this relationship was forced by the CARA population, without which
418 the relationship becomes non-significant (Fig 6b). No significant relationships between
419 the resilience indices and long-term survival were observed in Cabañeros (Fig. 6).

420

421 **Discussion**

422 This study made it possible to disentangle the effects of environment (i.e., site-to-site
423 plasticity), genetics (population differentiation) and G×E (i.e., population differences in
424 plasticity) on maritime pine responses in radial growth to two severe drought events.
425 Using two common garden trials, we were able to account for relevant factors that are
426 known to modulate tree responses to drought events, such as tree age, competitive status
427 and inter-tree competition (Fernández-de-Uña et al., 2015; Martín-Benito et al., 2008;
428 Martínez-Vilalta et al., 2012; Serra-Maluquer et al., 2018). These sources of variation
429 were discarded because we sampled dominant trees established simultaneously under
430 common environmental conditions according to a completely regular design. Inter-tree
431 competition was accounted for by incorporating an index of competition (surrounding
432 basal area to each experimental tree) as a covariate in the statistical models. Quantifying
433 plasticity and true genetic differentiation among populations is essential to examine the

434 adaptive value of drought-tolerance traits and strategies, and to suggest mitigation
435 management tools such as assisted migration (Aitken et al., 2016; Alberto et al., 2013).

436

437 *Drought response varies greatly across sites and drought events*

438 A first important finding of this study is that radial growth response to extreme drought
439 events largely depended on site. For each drought event, site was the most relevant
440 factor explaining variation in resilience components, even when environmental
441 differences between the test sites were not substantial (see Table ST1, Fig. 1).

442 Climate sensitivity and growth responses to drought have been shown to differ
443 across test sites in other conifer species, with differences attributed to contrasting
444 climate conditions across sites (McLane et al., 2011; Suvanto et al., 2016; Taeger et al.,
445 2013). In our study, however, the intensity, timing and duration of each extreme event
446 (1995 and 2005) were similar between sites (see Fig. 2). Thus, differences between sites
447 in response to extreme drought events should not be attributed to differences associated
448 to the extreme event itself, but to other environmental differences that must exist
449 between sites.

450 Environmental conditions in the two test sites were similar but not equal. These
451 differences were enough to imprint notable differences between sites in pine growth and
452 survival (Alia et al., 1997). Cabañeros was slightly wetter and colder than Riofrío, and
453 Riofrío has stonier and shallower soils, with higher slopes and irregular topography than
454 Cabañeros. The lower overall quality for pine development of the Riofrío test site may
455 have increased the impact of the drought events on all populations, as evidenced by the
456 greater reduction of radial growth rates during the event year and lower resistance to
457 drought at this site. Accordingly, pointer years coincided with the two drought events in
458 Riofrío, but only with the 2005 event in Cabañeros. Differences in edaphic and

459 geomorphological characteristics between sites may have particularly influenced early
460 seedling development affecting for example root development or biomass partitioning
461 (Chambel et al., 2007; de la Mata et al., 2014). This effect could have lasted to mature
462 ages and magnified after a drought event amplifying site differences in response to
463 climate stressors (Andivia et al., 2018). Little environmental differences triggering large
464 differences in how trees responded to climate extremes may have important
465 consequences when forecasting forest responses to climate change. Microenvironmental
466 factors such as topography, soil depth and stoniness –typically unexamined in previous
467 studies– may be much more relevant than previously recognized.

468 A second important finding of this study is that maritime pine responses to
469 drought were strongly dependent on the particular characteristics of the event. Previous
470 studies identified the intensity, timing and duration of the drought event as major
471 drivers of forest drought resilience components (Gazol et al., 2018; Taeger et al., 2013).
472 The two drought events studied here showed different intensity, timing and duration:
473 1995 was the last of six consecutive years of water deficit, while 2005 was a single very
474 dry year preceded and followed by several drought-free years (see Fig. 2a). These
475 differences likely explain the variable growth responses to the two drought events. For
476 example, the longer duration and higher intensity of the 1995 drought event is likely
477 behind the lower recovery observed for this event.

478 Inter-tree competition also influenced the resilience of maritime pine trees to the
479 extreme drought events as evidenced by the significant correlation with surrounding
480 basal area. According to previous findings (Linares et al., 2010; Martinez-Vilalta et al.,
481 2012), trees growing under stronger neighbouring competition were less resistant to
482 drought, probably because of reduced resource availability. However, the effect of
483 competition on recovery and relative resilience was positive. A possible explanation for

484 this counterintuitive result arises from the positive autocorrelation in growth that occurs
485 in experimental designs with multi-tree plots (16 trees, 4 rows \times 4 columns in our case)
486 (e.g. Zas, 2006). Besides microenvironmental effects, in multi-tree plot designs, most
487 trees that surround each experimental tree are from the same population and thus tend to
488 grow similar to the focal tree. This results in a positive autocorrelation which, based on
489 what was observed here, seems to surpass the negative autocorrelation between
490 competition and growth responses to drought. Previous studies also failed to detect
491 consistent and strong negative effects of competition on resilience components (Serra-
492 Maluquer et al., 2018).

493

494 *Patterns of growth response to drought among populations*

495 Sites and events not only influenced overall growth responses to drought and resilience
496 components, but also altered the variation observed among pine populations in their
497 drought resilience strategies (see significant interactions with populations in Table 1).
498 Patterns of among-population variation in recovery and relative resilience were
499 dependent on site, and variation in all resilience components except recovery was
500 dependent on the drought event. Repeatability across events in resilience components in
501 *Larix decidua* populations was low, reflecting a high event \times population interaction
502 (George et al., 2017). Similarly, in *Abies alba*, among-population variation in response
503 to different drought events varied greatly from one event to another, resulting in no
504 overall common patterns of variation (George et al., 2015). Viewed together, these
505 results highlight the importance of common-garden tests for exploring intraspecific
506 patterns of variation in drought responses and the need to analyse responses to multiple
507 events at different sites, irrespective of the environmental distance between them (de
508 Villedoreuil et al., 2016).

509 Despite the significant site \times population and event \times population interactions,
510 global general patterns of among-population variation were detected in growth
511 responses to drought and resilience components. Resistance and recovery in the face of
512 the extreme events showed a common population signal across sites and events. In
513 Riofrío, where the impact of the extreme events was more evident, population resistance
514 and recovery were strongly correlated across events. This is an important result when
515 inferring the adaptive value of the observed intraspecific variation. Maritime pine
516 populations from drier environments were less resistant to drought but recovered faster
517 than populations from mesic environments, a result that has been previously suggested
518 upon natural population surveys in maritime pine (Sánchez-Salguero et al., 2018), and
519 other tree species (Gazol et al., 2017; Stuart-Haentjens et al., 2018). Moreover, in the
520 site where the impact of drought was strongest (Riofrío), resistance and recovery were
521 negatively correlated, suggesting a trade-off between both components, as reported
522 elsewhere (Gazol et al., 2017; Zang et al., 2014).

523 These observations are consistent with results from physiological studies on young
524 pine trees, where the growth of populations from mild climates were less sensitive to
525 water deficit (i.e. they were more ‘resistant’ to drought) while populations from drier
526 conditions rapidly stop growing in response to water deficit (de la Mata et al., 2014;
527 Klein et al., 2014). This risky strategy of the Atlantic populations may lead to cavitation
528 and higher mortality rates under long-lasting or intense drought (Klein et al., 2013).
529 This is consistent with the negative correlation between population resistance and long-
530 term survival observed in Riofrío, where the intensity and impact of the drought events
531 were slightly higher. Moreover, higher plasticity in stomatal regulation (Feinard-
532 Duranceau et al., 2018) and/or greater capacity to accumulate carbohydrate reserves
533 (Suárez-Vidal et al., 2017) in populations from drought-prone environments may

534 explain their greater recovery capacity, as suggested in among-species studies (Gazol et
535 al., 2018). Recovery emerges as a key adaptive strategy for coping with severe water
536 deficit, and the significant positive correlation between recovery and long-term survival
537 (Fig. 6b) points into this direction. The relationship between recovery and survival may
538 be, however, overrated by values of recovery and survival of a single population
539 (CARA, from the driest region of the Iberian Peninsula, Fig. 6b). The outstanding
540 recovery of this population is likely due to the interruption of radial growth during the
541 year of the drought event, resulting in very large recovery estimates. As mentioned
542 before, minimizing growth in response to water deficit is a conservative strategy
543 typically shown by drought-adapted origins of Mediterranean pines (de la Mata et al.,
544 2014; Klein et al., 2014). Other pine species, however, may show contrasting
545 adaptations to drought, with trees adapted to stressful environments showing less
546 growth plasticity, and less plastic trees surviving better after extreme droughts (Ogle et
547 al., 2000). It should be noted that removing the CARA population from the analyses
548 makes the relationship between recovery and survival non-significant. The
549 interpretation of the recovery-survival relationship should be, thus, considered with
550 caution.

551

552 *Adaptive significance of growth resilience to drought*

553 Determining whether the observed patterns of intraspecific differentiation are
554 consequences of local adaptation processes is an important milestone for forecasting the
555 impact of climate change across the natural range of a given species and developing
556 management tools to mitigate its effect (de Villemereuil et al., 2016). As observed in
557 other conifer species such as *Pinus sylvestris* (Taeger et al., 2013), *Pseudotsuga*
558 *menziesii* (Montwé et al., 2015), or *Larix decidua* (George et al., 2017), we found

559 intraspecific genetic variation in drought resilience components among the studied
560 populations. Particularly, resistance to and recovery from the extreme drought events
561 showed consistent among-population variation. Because the main drought coping
562 strategies of a given species tend to be those that are differentiated among populations
563 as a consequence of local adaptation to environmental heterogeneity (Zang et al., 2014),
564 we hypothesize that both resistance and recovery are key adaptive drought-tolerance
565 strategies in this Mediterranean species.

566 Differentiation among populations can arise not only by divergent selection
567 across a heterogeneous environment, but also by neutral processes like drift, associated
568 with the demographic history of the species (George et al., 2017). Marked and
569 structured neutral differentiation can be especially important in Palearctic species that
570 suffered strong genetic bottlenecks during the last ice age, which is the case of maritime
571 pine (Jaramillo-Correa et al., 2015). Neutral and adaptive variation could be
572 disentangled by gathering information on neutral genotypic markers and using
573 population genetics approaches such as F_{st} - Q_{st} comparisons (Alcaide et al., 2019), or
574 accounting for neutral variation in the statistical model fitting (López-Goldar et al.,
575 2019). Although we lack genotyping information of our material, several of our results
576 suggest that the patterns of variation observed are the result of adaptive processes. On
577 the one hand, the conserved patterns across sites and drought events of among-
578 population variation in resistance to and recovery from the extreme climate events
579 points to a marked genetically-driven pattern of variation. On the other hand, the
580 consistent and biologically meaningful relations between variation in resilience
581 components and climate conditions at origin and the relationships between population
582 resistance and population recovery with long-term survival suggest that the genetic
583 variation in these resilience indices has an adaptive origin. The later relationships

584 suggest that the resilience components may be relevant for tree fitness, and could
585 therefore be subjected to selective pressures associated with extreme drought events.
586 However, several limitations of these relationships must be recognized. First, survival
587 was evaluated at the time of core sampling (when trees were 43 years old), and it can't
588 be assumed that the cumulated mortality of trees to this date was totally related to the
589 occurrence of the two extreme drought events. Early mortality during the first years
590 after establishment could, for example, be disturbing the biological meaning of these
591 relationships. Second, the relationship between recovery and survival was overridden by
592 the outstanding values of one population (see discussion above), diminishing the
593 reliability of the relationship. Finally, the different population survival rates may have
594 generated different competition levels among populations which, in turn, may have
595 affected tree responses to the drought events. This potential bias was, however,
596 accounted for by including inter-tree competition in the statistical models and adjusting
597 population least square means for variation in competition among trees. Altogether,
598 although providing relevant information, the relationships between the resilience
599 components and survival must be managed with caution. Further studies are needed to
600 confirm the adaptive value of the resilience components in response to extreme drought
601 events.

602 Other studies reported evidence that the components of resilience to extreme
603 climate events assessed here are integrative traits reflecting drought-coping adaptations.
604 For example, resilience components have been associated with functional molecular
605 markers related to drought (Heer et al., 2018; Housset et al., 2018; Trujillo-Moya et al.,
606 2018), and several studies have shown marked relations between among-population
607 variation in resilience components and climate conditions at origin (Housset et al.,
608 2018; Stuart-Haentjens et al., 2018; Trujillo-Moya et al., 2018). Moreover, repeatability

609 of resilience components across sites and events has been found to vary significantly
610 among populations (George et al., 2017), and this has been interpreted as a sign of
611 adaptation, given that populations subjected to strong selective pressure (i.e., inhabiting
612 drought-prone environments) showed lower within-population variation and higher
613 repeatability across sites and environments (George et al., 2017).

614

615 **Data Accessibility Statement**

616 BAI chronology data are currently being submitted to DIGITAL.CSIC repository, and
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618

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632

633 **Conflict of interest**

634 The authors declare no conflict of interest.

635

636 *Author contributions:* RZ , VR and LS conceived the ideas; RA contributed with
637 previous data from the field trials; all authors participated in field sampling; VR
638 processed the samples and did the dendrochronology assessments; RZ and VR analysed
639 the data; RZ led the writing with assistance from all authors.

640

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864 **Table 1.** Main effects of test Site (Cabañeros and Riofrío), Population (10 levels) and
865 drought Event (1995 and 2005) and their interactions on the drought resilience
866 components, defined as resistance, recovery, resilience and relative resilience (Lloret et
867 al., 2011). Effect of the drought event and the corresponding interactions were analysed
868 as within-subject effects. Inter-tree competition (estimated as the sum of the basal area
869 of neighbouring living trees) was included as a covariate in the model and, in case of a
870 significant effect, the sign of the covariation indicated within brackets. Degrees of
871 freedom (DF), F ratios and associated probability levels are shown. Significant effects
872 ($p < 0.05$) are shown in bold.

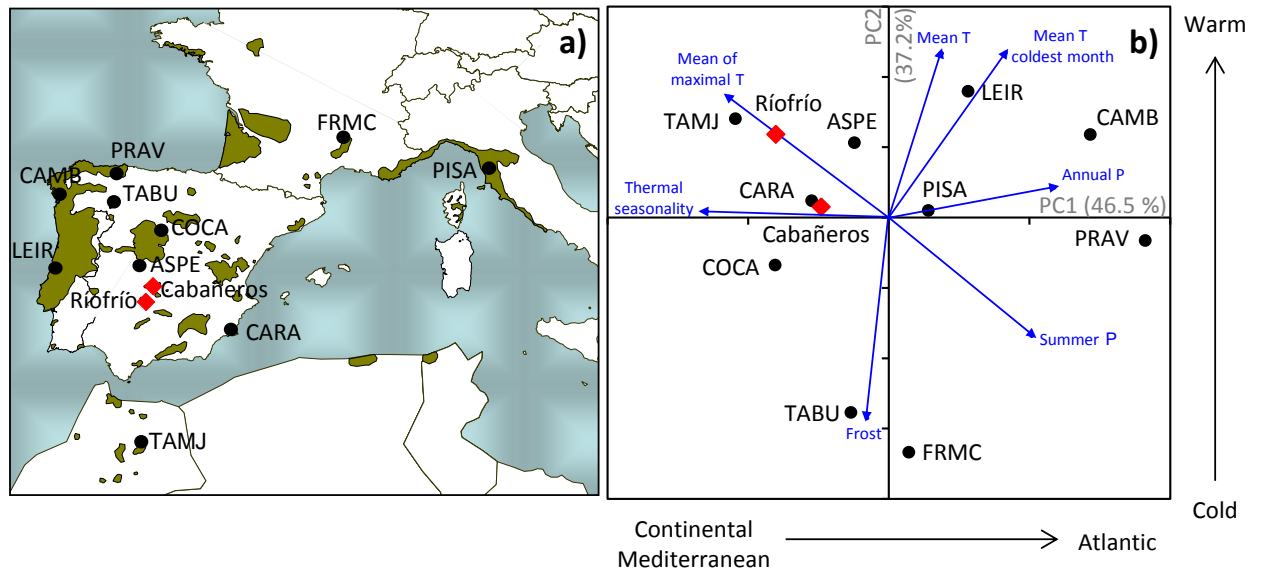
	DF	Resistance		Recovery		Resilience		Relative resilience	
		F	P>F	F	P>F	F	P>F	F	P>F
<i>Across subjects</i>									
Site	1, 6	138.8	<0.001	119.2	<0.001	3.5	0.111	127.4	<0.001
Population	9, 44	4.0	0.001	3.2	0.005	1.4	0.223	0.7	0.723
Site x Pop	9, 44	1.5	0.198	4.3	<0.001	1.4	0.237	2.4	0.024
<i>Within subjects</i>									
Event	1, 252	84.6	<0.001	16.0	<0.001	13.9	<0.001	105.7	<0.001
Event x Site	1, 252	115.4	<0.001	0.1	0.818	2.5	0.116	48.9	<0.001
Event x Pop	9, 252	4.7	<0.001	0.3	0.961	3.9	<0.001	5.6	<0.001
Event x Site x Pop	9, 252	3.6	<0.001	1.2	0.277	2.2	0.023	2.6	0.008
Neighbouring competition	1, 252	10.9	0.001 (-)	15.5	0.000 (+)	0.0	0.891	5.1	0.024 (+)

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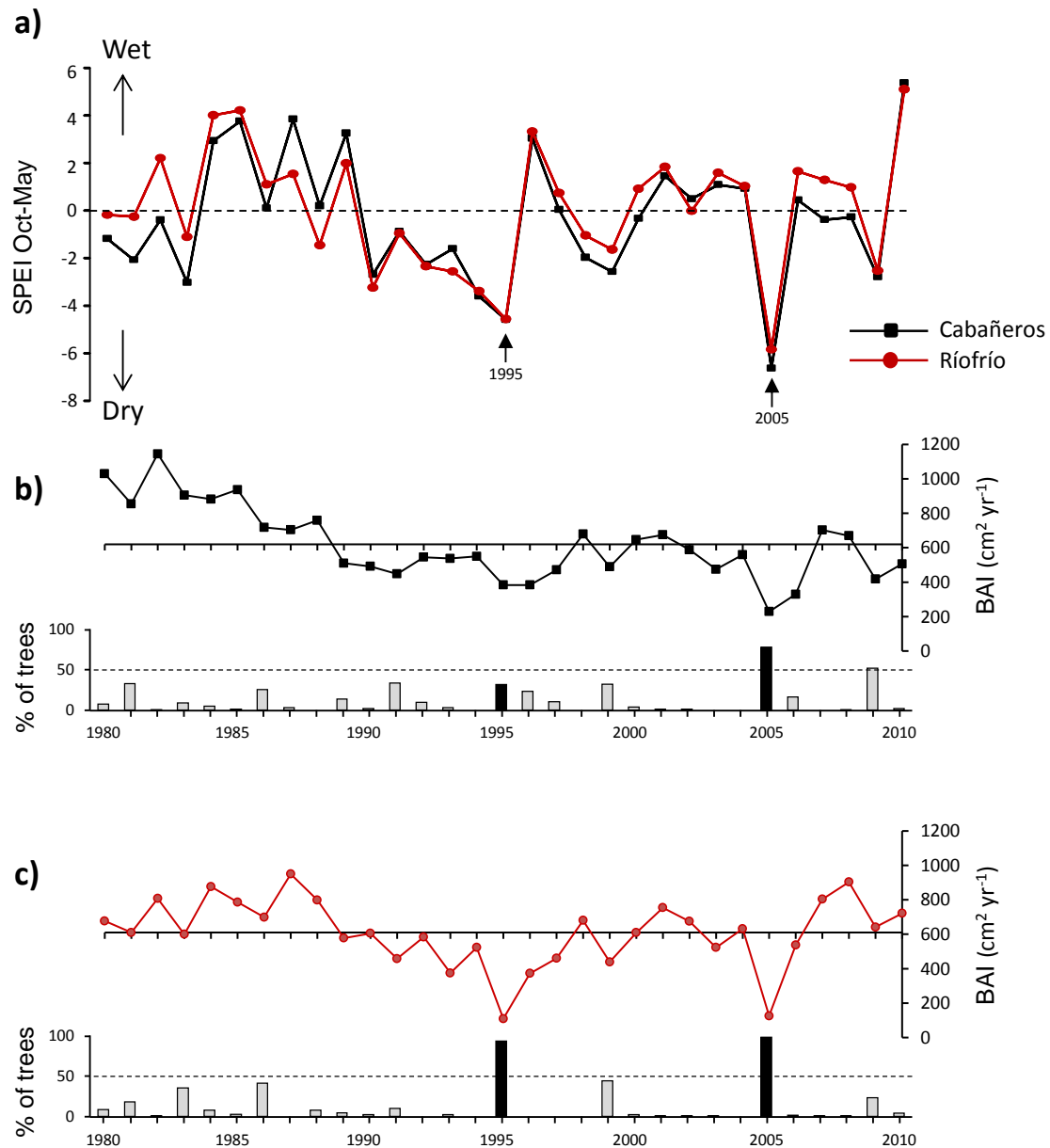


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878 **Fig. 1.** a) Location of the two *Pinus pinaster* genetic test sites (red diamonds) in central
 879 Spain and origin of the 10 populations (black dots) across the species range of
 880 distribution [green area, EUFORGEN 2009, www.euforgen.org]. See Table ST1 for
 881 population abbreviations. b) Principal component analysis of climate data showing the
 882 loadings of the climate variables on the first two principal components (blue arrows)
 883 and the spread of the 10 maritime pine populations (black dots) and the test sites (red
 884 diamonds) along the resulting bi-dimensional space.

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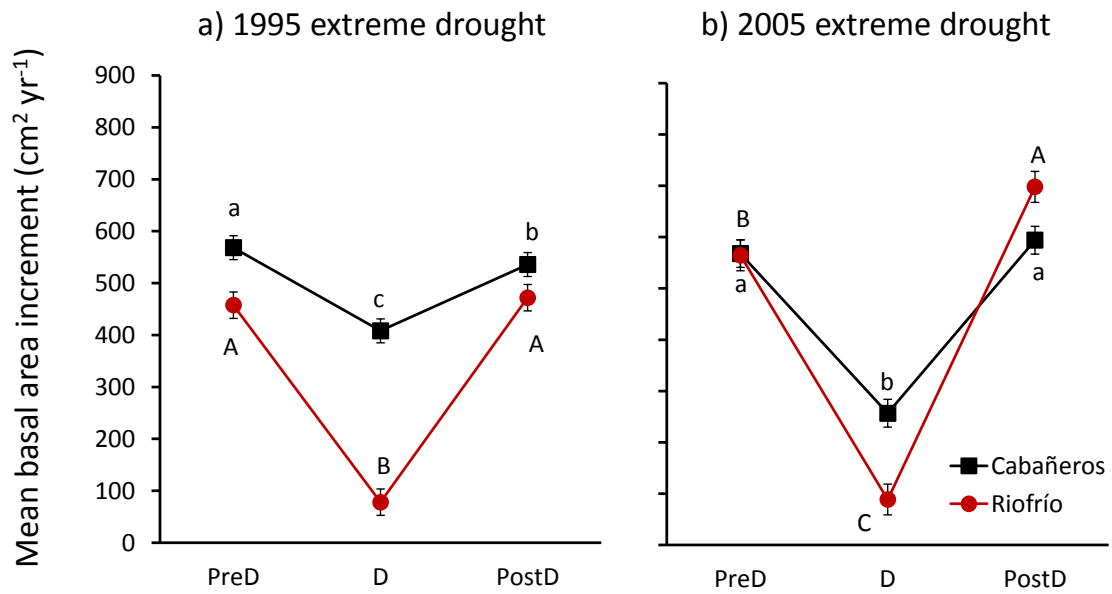
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888 **Fig. 2.** Annual variation in the standardised precipitation-evapotranspiration index
 889 (SPEI) of the wet period of the year in the area (October to May) (a), and annual
 890 patterns basal area increment (BAI) and negative pointer years at the Cabañeros (b) and
 891 Ríofrío (c) test sites. The extreme negative values of SPEI in 1995 and 2005 are
 892 indicated by arrows. Negative pointer years (bars) are expressed as a percentage of trees
 893 showing a sharp growth reduction ($\geq 50\%$ BAI reduction) in comparison with growth in
 894 the previous three years, according to the *dplR* package (Bunn, 2010). Black bars
 895 indicate the extreme drought events of 1995 and 2005 considered in this work.

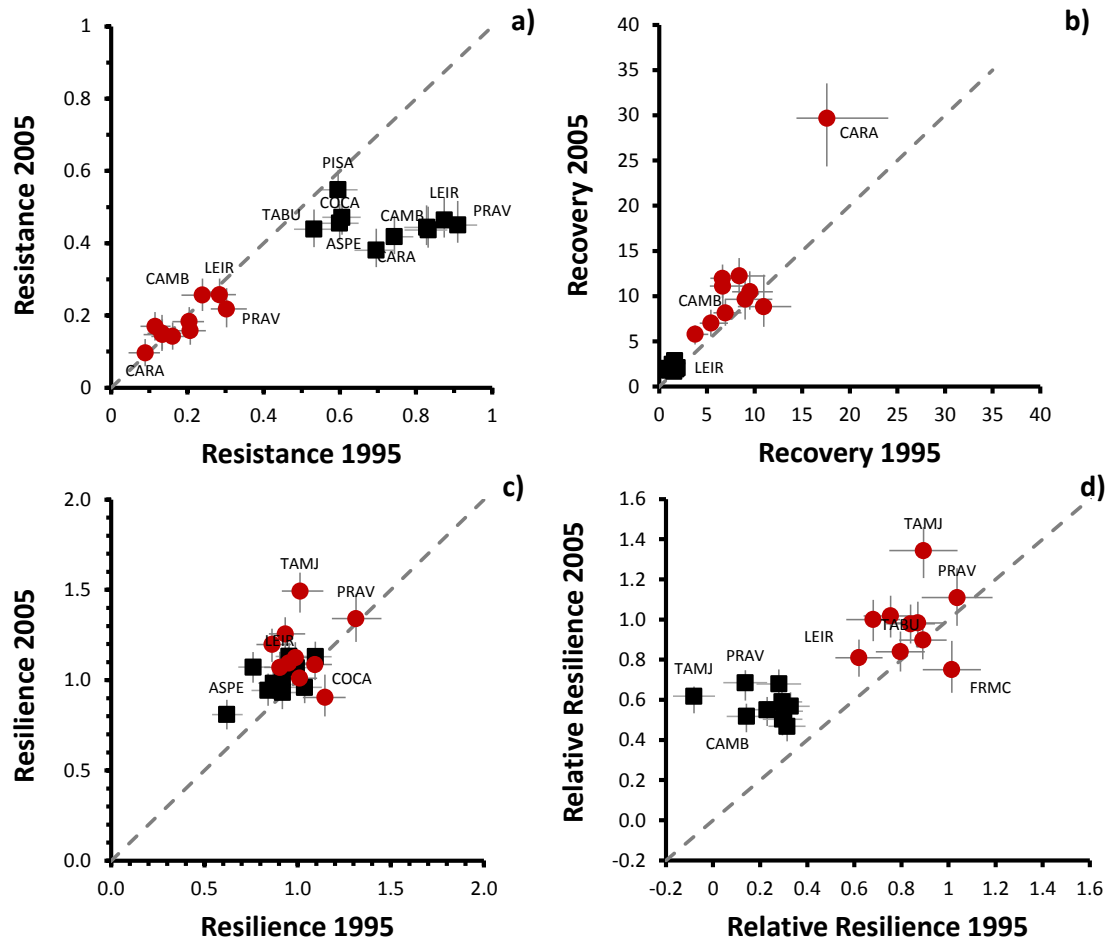
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899 **Fig. 3.** Least square means (\pm s.e.) of basal area increment three years before (PreD), the
 900 year during (D) and three years after (PostD) the drought events of 1995 (a) and 2005
 901 (b) events at Cabañeros (black squares) and Riofrío (red dots) test sites. Growth during
 902 the three periods were significantly different for both events ($F_{2,507} = 396.7$, $p < 0.001$
 903 for the 1995 drought; $F_{2,507} = 564.8$, $p < 0.001$ for the 2005 drought) but differences
 904 varied depending on the test site (drought \times site effect, $F_{2,507} = 85.3$, $p < 0.001$ for the
 905 1995 drought; $F_{2,507} = 41.3$, $p < 0.001$ for the 2005 drought). Different lowercase and
 906 uppercase letters indicate significant differences (Tukey's HSD post-hoc test) between
 907 the three periods compared at Cabañeros and Riofrío, respectively.

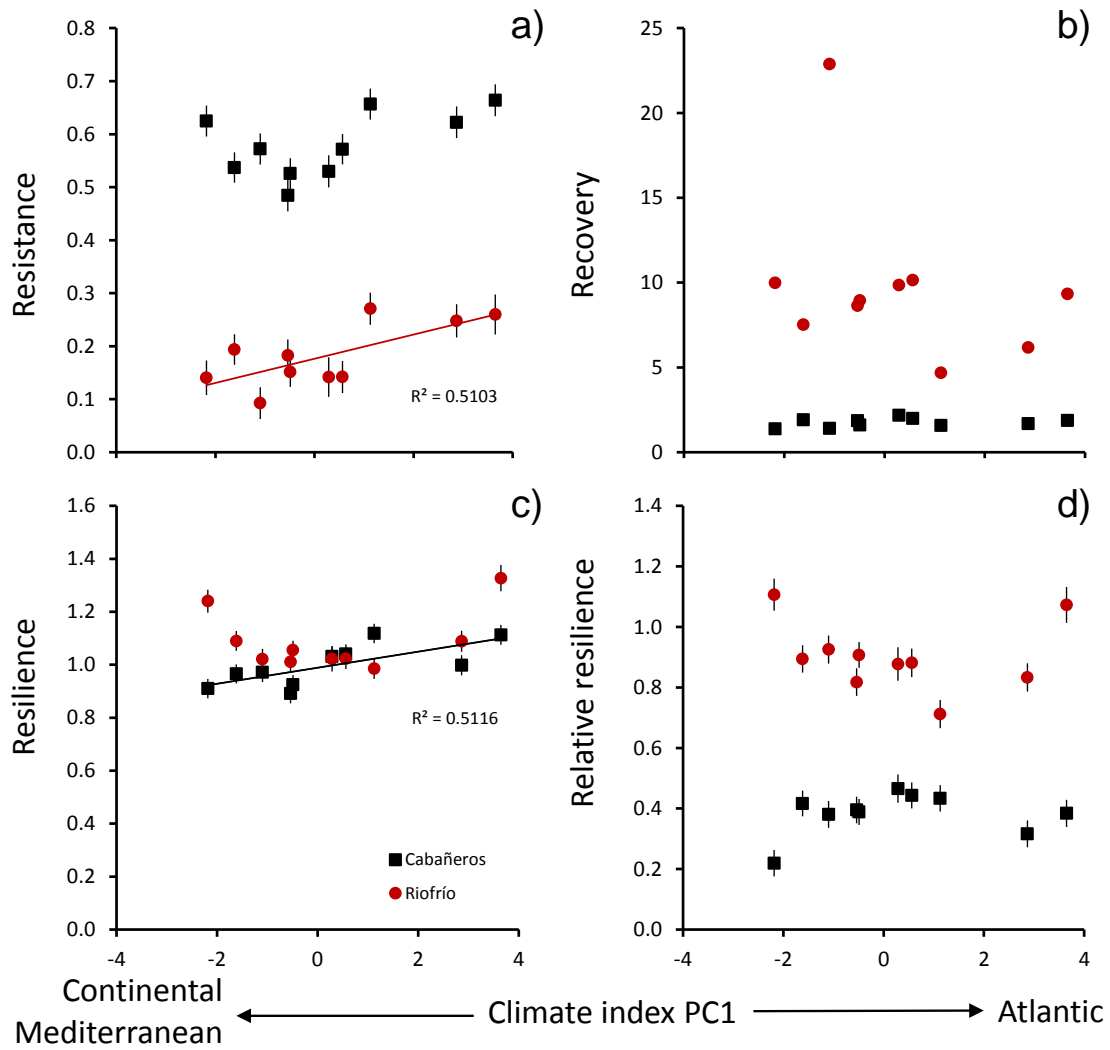
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910 **Fig. 4.** Components of tree resilience in response to the extreme drought events of 1995
 911 (horizontal axes) and 2005 (vertical axes) for 10 maritime pine populations established
 912 at two test sites (Cabañeros: black squares; Riofrío: red dots). Least square means \pm
 913 standard errors are shown for (a) resistance, (b) recovery, (c) resilience and (d) relative
 914 resilience. The diagonal dashed line represents the 1:1 relation between the drought
 915 events. See Table ST1 for population abbreviations.

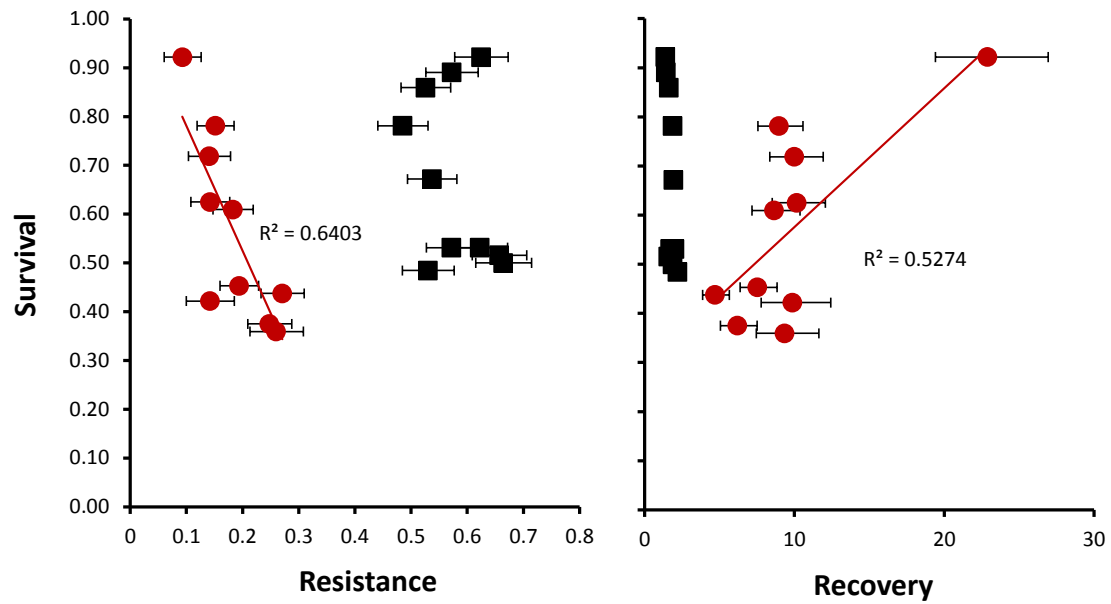
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918 **Fig. 5.** Relation between climate conditions at origin (PC1 climate index) of the 10 *P.*
 919 *pinaster* populations in Cabañeros (black squares) and Riofrío (red dots) test sites and
 920 the corresponding population least square means for the resistance (a), recovery (b),
 921 resilience (c), and relative resilience (d) calculated across the extreme drought events
 922 studied (1995 and 2005). Linear trends and the corresponding r^2 values are shown for
 923 significant relations ($p < 0.05$).

924



925

926 **Fig. 6.** Relations between (a) resistance and (b) recovery of the 10 maritime pine
 927 populations and overall provenance survival at age 43 at the Riofrío (red circles) and
 928 Cabañeros (black squares) test sites. Linear trends and the corresponding r^2 values are
 929 shown ($p < 0.05$). Dots are the least square means \pm standard errors across the two
 930 drought events, 1995 and 2005.

931

Supplementary material

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934 Dendroecology in common gardens: Population differentiation and plasticity in
935 resistance, recovery and resilience to extreme drought events in *Pinus pinaster*

936

937 Rafael Zas, Luis Sampedro, Alejandro Solla, María Vivas, María J. Lombardero,
 938 Ricardo Alía, Vicente Rozas

939

940 Methods SM1. Estimating monthly climatic data in the two test sites during the
941 studied period (1980-2010)

942 Monthly temperature series (mean, maximum, minimum, mean of maximums and mean
 943 of minimums) and precipitation series (monthly precipitation and number of rainy days
 944 per month) in the two test sites were estimated upon data of different climatic stations of
 945 the National Agency of Meteorology (AEMET, Ministerio para la Transición
 946 Ecológica, <http://www.aemet.es>) situated in the proximities of the two test sites.

947 Temperature series in Cabañeros and Riofrío were estimated using data from 4 and 3
 948 nearby stations, respectively (Table SM1), while precipitation series were estimated
 949 upon the 2 closest climatic stations of each test site (Table SM1). Stations were selected
 950 based not only in their proximity but also in the correspondence between the altitude
 951 and the relative geomorphological position in relation to the test sites.

952 Climatic stations (AEMET, <http://www.aemet.es>) used for the estimation of monthly
 953 temperature (T) and precipitation (P) series in the two test sites (CAB: Cabañeros, and
 954 RIO: Riofrío).

Station Code	Station Name	Altitude (m)	Long	Lat	Variables	Site
3298X	San Pablo de los Montes	917	383957	4378346	T	CAB
4123	Los Cortijos de arriba	775	407846	4352300	T	CAB
4193Y	El Robledo (Automática)	580	389109	4341690	T	CAB/RIO
4202E	Piedrabuena (El Casarejo)	660	382783	4330928	T	RIO
4316	Tamurejo	550	331393	4317335	T	RIO
4189	Pantano de Torre Abraham	697	391809	4358241	T	CAB
4171I	Retuerta del Bullaque	760	382188	4364610	P	CAB
4189	Pantano de Torre de Abraham	697	391809	4358241	P	CAB
4224	Arroba de los Montes (1)	615	366706	4334761	P	RIO
4224A	Arroba de los Montes (2)	615	366588	4334856	P	RIO

955

956 Monthly temperatures were obtained as the inverse distance-weighted means of
957 the temperatures of the selected climatic stations for each site, applying a correction
958 factor of $-0.005\text{ }^{\circ}\text{C}$ per meter of altitude difference between the climatic station and the
959 test site. Missing data in the original stations was previously predicted by regression
960 models fitted between the available temperatures in the climatic station and the CRU TS
961 4.01 data set (Mitchell and Jones, 2005). CRU climate time series for the period 1901–
962 2016 were taken from the Climate Explorer of the Royal Netherlands Meteorological
963 Institute (<http://climexp.knmi.nl/>) for the 0.5° longitude \times 0.5° latitude sector in which
964 every climatic station was located.

965 Monthly precipitations in each site were estimated directly from climatic stations
966 that were very close to the test sites. In the case of Riofrío, precipitations were obtained
967 as the mean value of the data of the two closest stations, both sited $<2\text{ Km}$ from the test
968 site. In Cabañeros, data from the closest climatic station ($<5\text{ km}$ from the test site) was
969 largely incomplete. Missing data in this station was completed by regression models
970 fitted between the precipitation series in this climatic station and those in the next closer
971 station (12.5 km from Cabañeros). Additionally, remaining missing data in the two test
972 sites were filled upon CRU TS 4.01 data using a similar procedure as that describer for
973 temperatures.

974

975 **References**

976 Mitchell, T.D., Jones, P.D. 2005. An improved method of constructing a database of
977 monthly climate observations and associated high resolution grids. *International*
978 *Journal of Climatology* 25, 693–712.

979

Dendroecology in common gardens: Population differentiation and plasticity in resistance, recovery and resilience to extreme drought events in *Pinus pinaster*

Rafael Zas, Luis Sampedro, Alejandro Solla, María Vivas, María J. Lombardero, Ricardo Alía, Vicente Rozas

Methods SM2. Statistical mixed model for the analysis of growth responses to the extreme drought events

Growth responses to each extreme drought event were analysed separately by repeated mixed model analysis applied to tree-level BAI data. The following statistical mixed model was fitted (random effects in italics):

$$BAI = \mu + S + POP + POP \times S + Dr + Dr \times S + Dr \times POP + Dr \times POP \times S + B(S) + B \times POP(S) + Tree(B \times POP \times S) + NBA + \varepsilon$$

where μ is the overall mean basal area increment, S, POP, and POP×S stand for the across-subject fixed effects of the site, the population and their interaction, Dr represent the within-subject fixed effect (repeated measures on the same subjects, the trees) of the drought factor (PreD, D and PostD), B(S) and B×POP(S) represent the random variation between blocks within sites and between whole plots of the multi-tree block design, respectively, *Tree(B×POP×S)* is a random factor identifying each tree within each block, population and site and was included in the model to account for the autocorrelation among the repeated measures, NBA is a continuous fixed covariate accounting for the basal area surrounding each tree ten years before each climatic event as a proxy of competition, and ε is the random residual.

The across-subject effects of populations, sites and their interaction accounts for global differences among populations and sites in mean annual radial growth, and has nothing to do in relation to the extreme drought events.

Responses to the drought events are tested by the within-subject factors. The drought factor accounts for the plastic response in BAI to the drought event (temporal plasticity) while the interactions between this factor and the sites and the populations represent (spatial) plasticity to environmental variation between sites and among-population variation in these responses to drought, respectively.

Supplementary material

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1014 **Dendroecology in common gardens: Population differentiation and plasticity in**
 1015 **resistance, recovery and resilience to extreme drought events in *Pinus pinaster***

1016

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1020 **Table ST1.** Characteristics of the ten *Pinus pinaster* provenances at their origin and the
 1021 two experimental test sites. Average population and site survival and growth are also
 1022 shown.

Site name	Code	Genetic group ¹	Elevation (m)	Latitude (°)	Longitude (°)	Annual precipitation (mm)	Mean annual temperature (°C)	Survival (%) ⁴	Total height (m) ⁵	Diameter (cm) ⁵
<i>Populations</i>										
Arenas	ASPE	Central Spain	634	40.209	-5.058	1217 ²	13.7	85.9	11.2	20.4
Tabuyo	TABU	Central Spain	1019	42.314	-6.220	734 ²	9.7	78.1	9.6	17.9
Pravia	PRAV	Northern Spain	146	43.555	-6.211	1299 ²	13.3	50.0	10.6	18.4
Cambados	CAMB	Northern Spain	28	42.526	-8.770	1550 ²	14.7	53.1	10.4	17.4
Coca	COCA	Central Spain	783	41.248	-4.493	454 ²	12.3	67.2	9.8	19.4
Pisa	PISA	Northern Italy	100	43.717	10.384	860 ³	13.9	53.1	10.0	17.7
Massif Central	FRMC	Atlantic France	400	45.001	3.75	839 ³	13.1	48.4	11.4	17.5
Tamjout	TAMJ	Western Africa	1600	33.867	-4.033	336 ³	14.2	92.2	10.3	18.6
Leiría	LEIR	Atlantic	60	39.751	-8.917	934 ³	15.6	51.6	11.5	19.3
Caravaca	CARA	Eastern Spain	1063	38.145	-1.984	481 ²	13.0	89.1	10.9	19.8
<i>Test sites</i>										
Cabañeros			1023	39.396	-4.475	796	12.4	65.9	11.0	18.8
Riofrío			725	39.134	-4.536	719	14.2	57.0	10.1	18.5

1023

1024 ¹ Bucci et al. (2007)

1025 ² Data from CRU TS 4.01 model for the period 1950-2000

1026 ³ Data from the model of Gonzalo (2007) for the period 1950-2000

1027 ⁴ Survival at the time of sampling (age 43)

1028 ⁵ Diameter at breast height and total tree height at age 32 (data recalculated from Alía et al.,
 1029 2001)

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Supplementary material

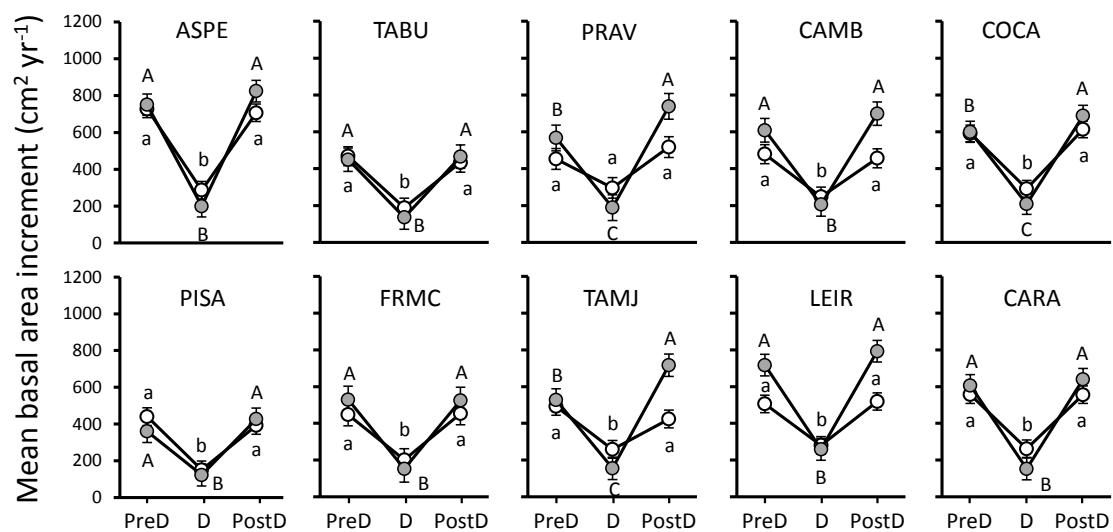
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1033 **Dendroecology in common gardens: Population differentiation and plasticity in** 1034 **resistance, recovery and resilience to extreme drought events in *Pinus pinaster***

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1038

1039 **Fig. SF1.** Least square means (\pm s.e.) of basal area increment three years before (PreD),
 1040 the year during (D) and three years after (PostD) the drought events of 1995 (white
 1041 dots) and 2005 (grey dots) of the ten studied populations averaged across the two test
 1042 sites. Growth during the three periods were significantly different for both events ($F_{2,507}$
 1043 = 396.7, $p < 0.001$ for the 1995 drought; $F_{2,507} = 564.8$, $p < 0.001$ for the 2005 drought)
 1044 but differences varied depending on the population (drought \times population interaction,
 1045 $F_{18,507} = 3.6$, $p < 0.001$ for the 1995 drought; $F_{18,507} = 3.3$, $p < 0.001$ for the 2005
 1046 drought). See Table ST1 for abbreviations of provenances. For each population,
 1047 different letters indicate significant differences between the three compared periods
 1048 based on Tukey's HSD post-hoc test (lowercase letters for the 1995 event and uppercase
 1049 for the 2005 event).

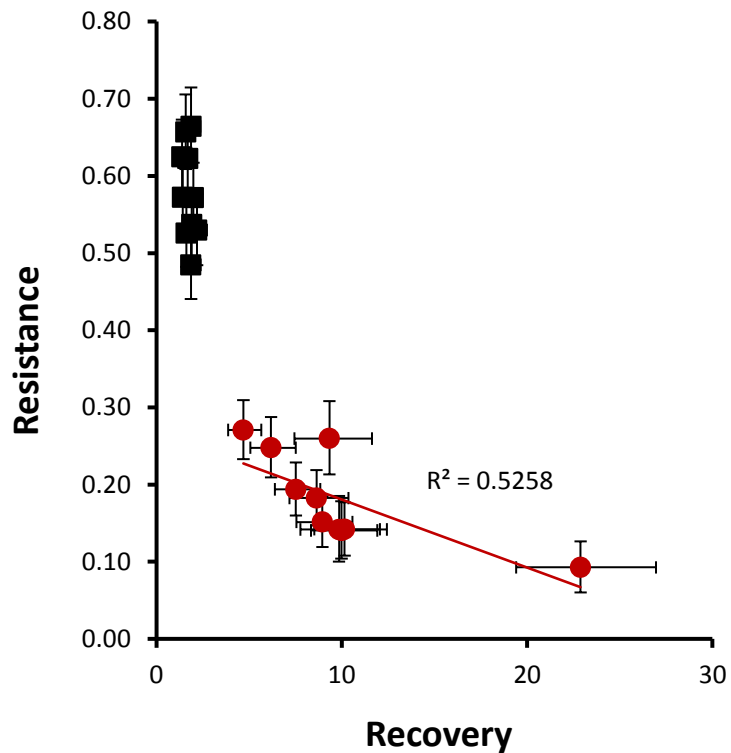
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Supplementary material

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Dendroecology in common gardens: Population differentiation and plasticity in resistance, recovery and resilience to extreme drought events in *Pinus pinaster*

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Fig. SF2. Relationship between resistance to and recovery from extreme drought events across 10 maritime pine populations established at two test sites (Cabañeros: black squares; Riofrío: red dots). Least square mean \pm standard errors (polled across the two drought events, 1995, 2005) are shown. For significant relationships, linear trend and the corresponding r^2 value are shown.

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Supplementary material

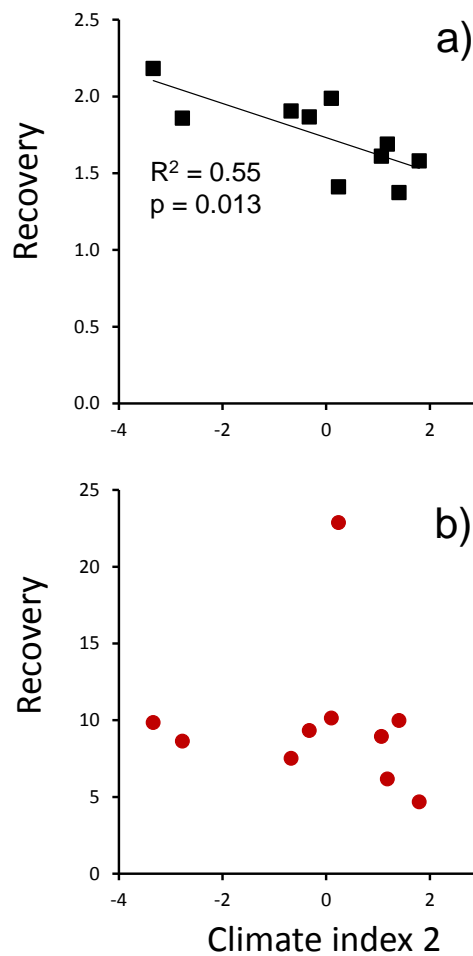
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1070 **Dendroecology in common gardens: Population differentiation and plasticity in**
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1074 Ricardo Alía, Vicente Rozas



1075

1076 **Fig. SF3.** Relation between climate index 2 (a proxy of thermal conditions) at origin of
1077 the 10 *P. pinaster* populations and the corresponding population least square means for
1078 growth Recovery after the two extreme drought events (1995 and 2005) in Cabañeros
1079 (a, black squares) and Riofrío (b, red dots) test sites. Linear trends and the
1080 corresponding r^2 values are shown for significant relations ($p < 0.05$).

Figure 1

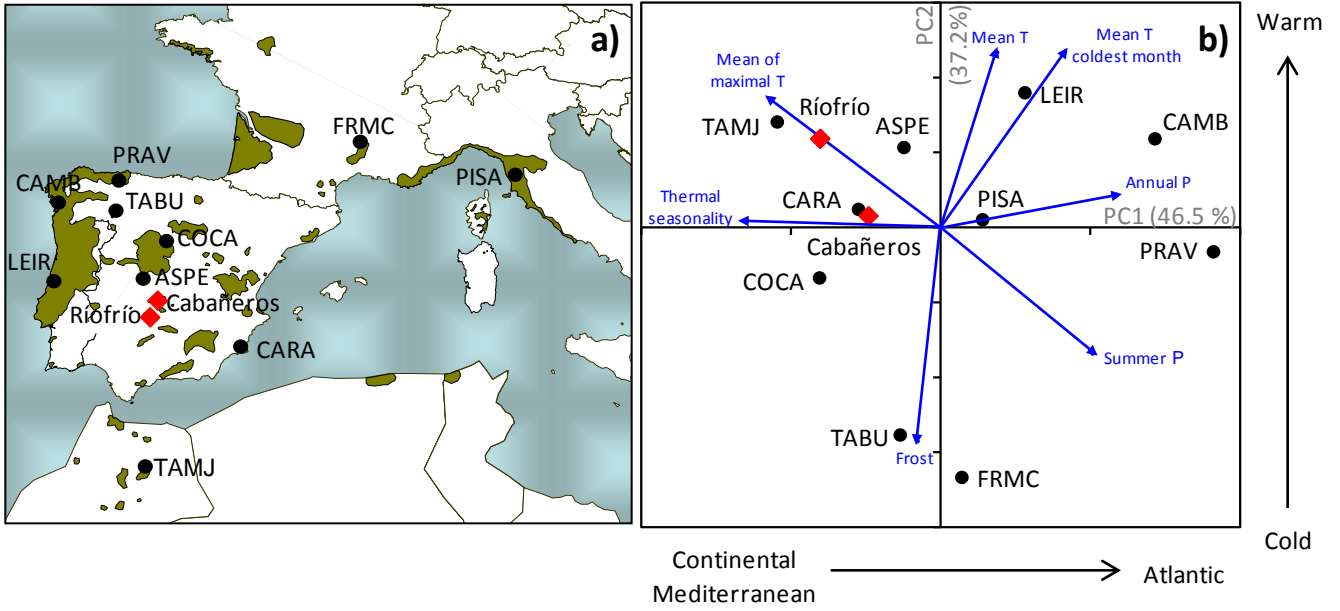


Figure 1

Figure2

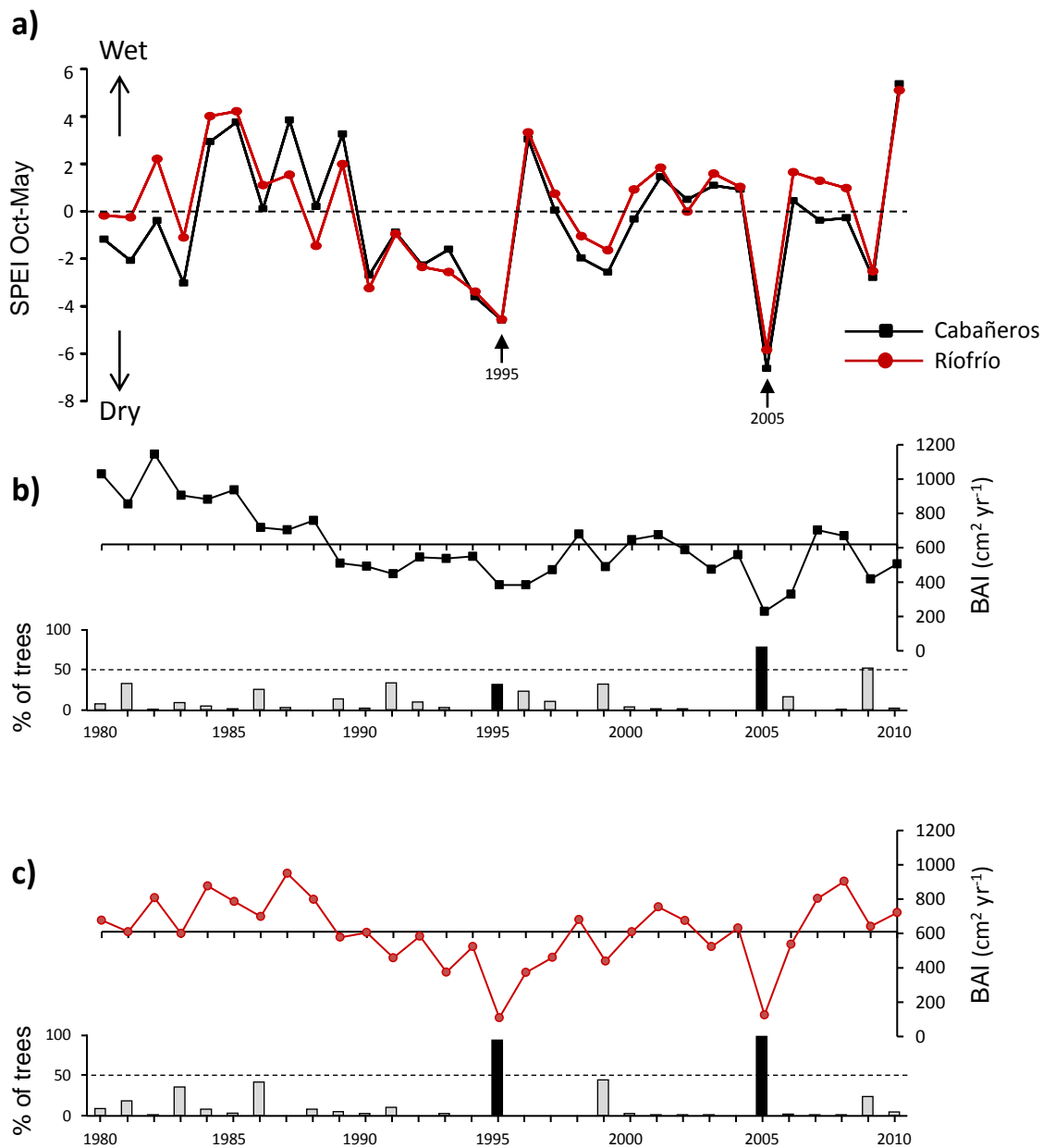


Figure 2

Figure3

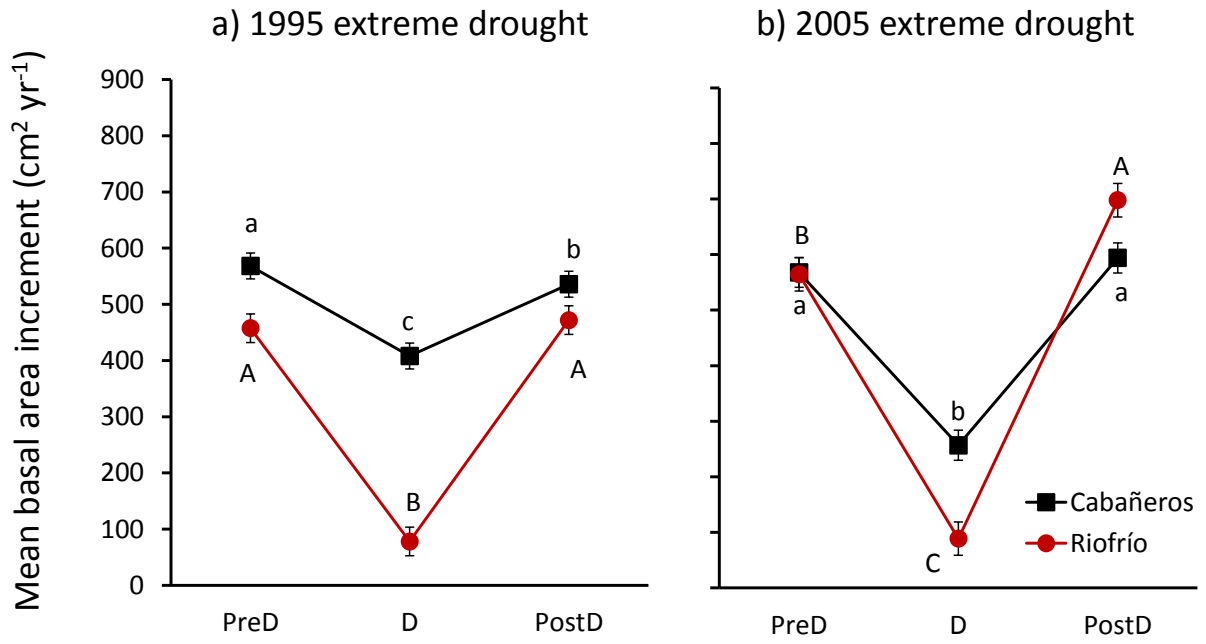


Figure 3

Figure4

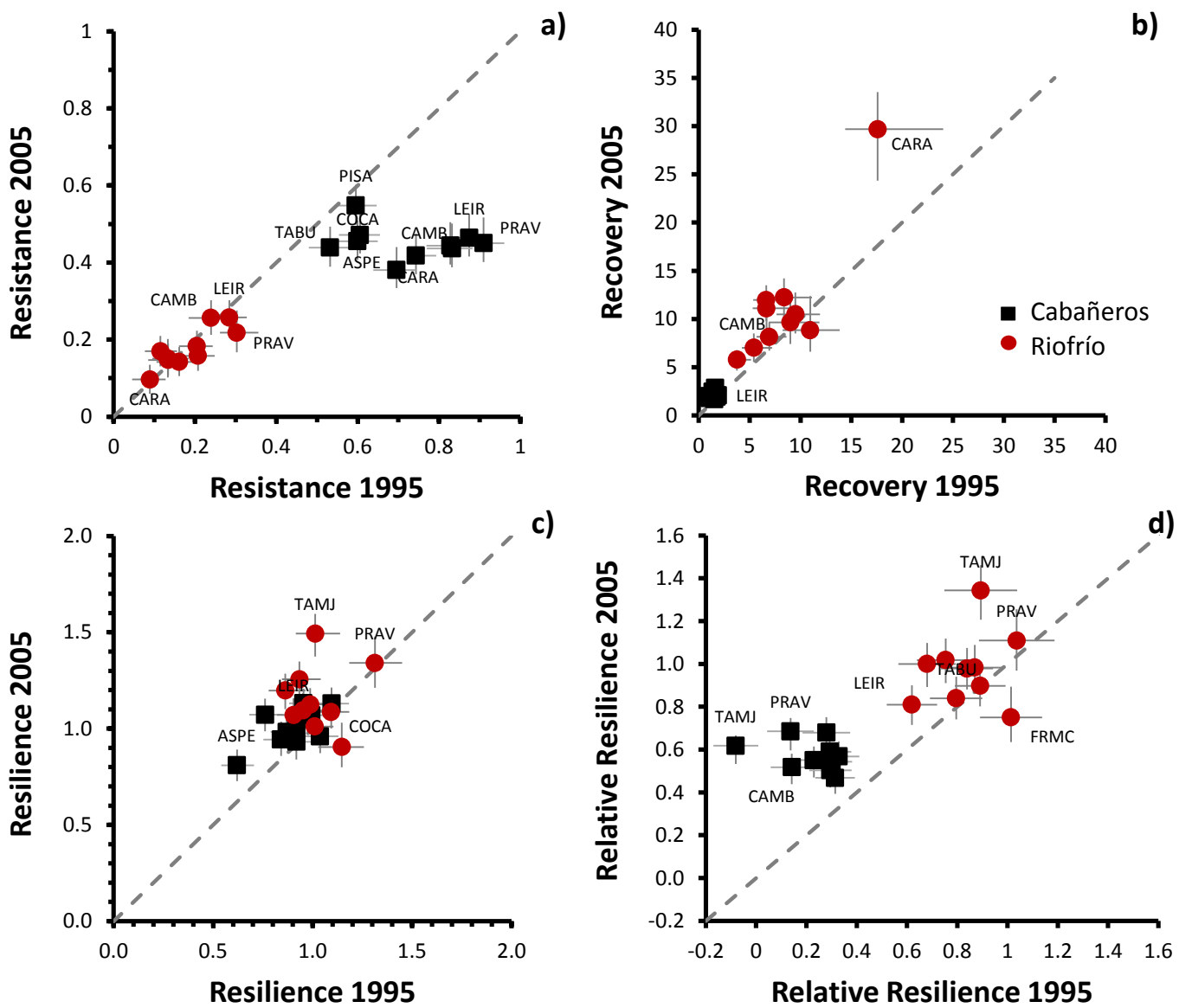


Figure 4

Figure5

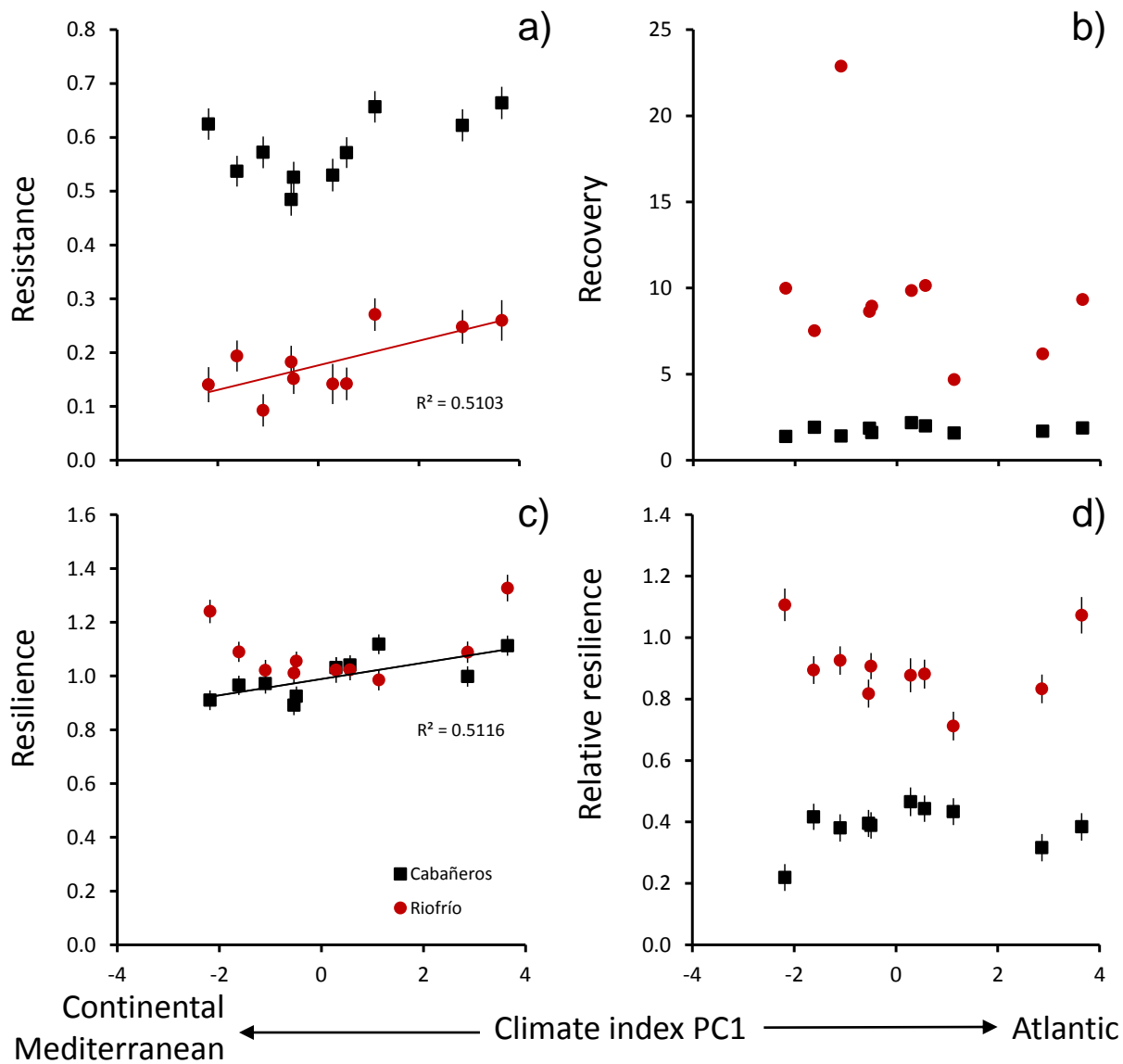


Figure 5

Figure6

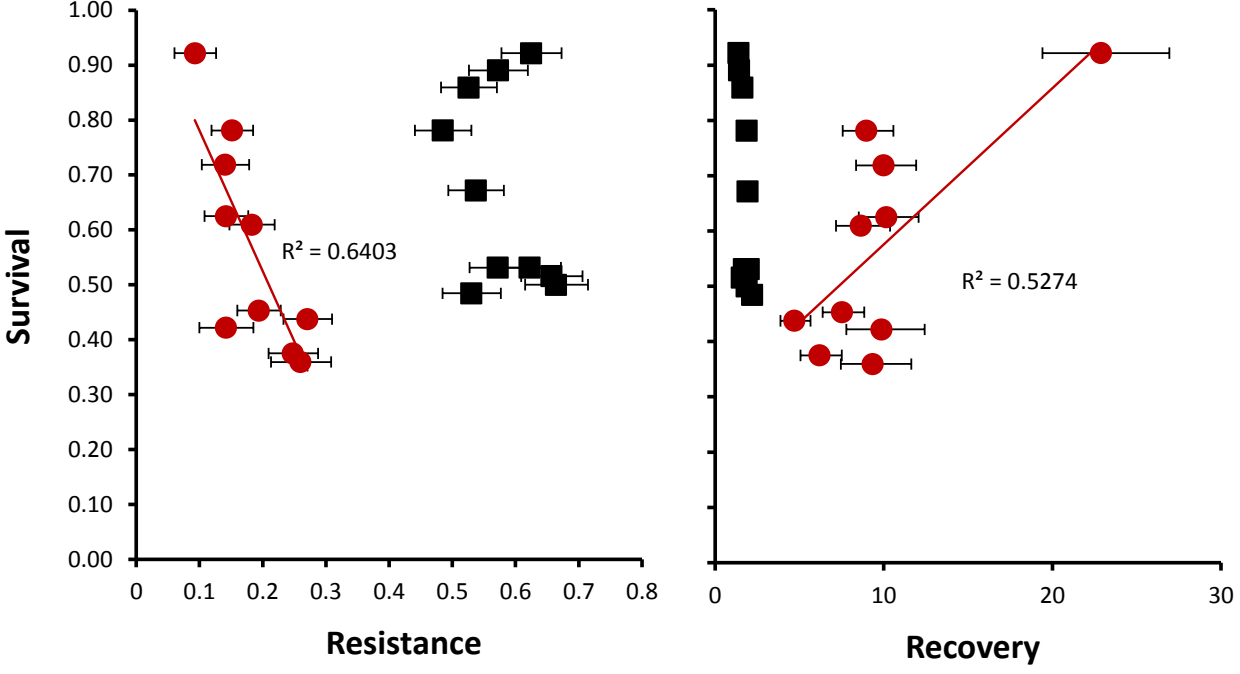


Figure 6

e-component

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